



Script generated by TTT

Title: profile1 (18.06.2013)
 Date: Tue Jun 18 11:59:08 CEST 2013
 Duration: 90:22 min
 Pages: 65

Data Mining: Metric Clustering

Finding Clusters in Profiles

Examples for profile elements that can be embedded in metric spaces:

- **Location & Velocity:** Metric space: $(\mathbb{R}^3, \|\cdot\|)$
- **Text** describing Interests: Metric space: $(\mathbb{R}^{|\text{Voc}|}, \|\cdot\|)$ where Voc denotes the Vocabulary of the text.

"I like to dance samba,
 bake pizza, watch tv and
 plant trees in the garden. I
 also like to bake cakes."



I	2
like	2
to	2
dance	1
samba	1
bake	2
pizza	1
watch	1
tv	1
and	1
plant	1
trees	1
in	1
the	1
garden	1
also	1
cakes	1

Often: Instead of
 term-frequency
 (tf) alone: use
 term-frequency *
 inverse document
 frequency (idf);
 $idf = \log(\#of\ docs / \#of\ docs)$



Finding Clusters in Profiles

- How do we compute **clusters in metric spaces?**
- **Group models:** How do we compute **socially meaningful clusters** in metric spaces (and thus avoid quasi-groups)?
- First some **notations / basics:**
 - In graph clustering we had: A graph clustering $\mathbf{C}=\{C_1, C_2, \dots, C_K\}$ is a partition of V into non-empty subsets C_k
 - Now: **clustering** $\mathcal{C} : \mathcal{X} \rightarrow \mathcal{I}$: mapping of a metric value space \mathcal{X} to a set of cluster indices \mathcal{I}
 - Clusterings can be:
 - **exclusive or non-exclusive**
 - **crisp or fuzzy**
 - **hierarchical or non-hierarchical**



Finding Clusters in Profiles

- How do we compute **clusters in metric spaces**?
- **Group models**: How do we compute **socially meaningful clusters** in metric spaces (and thus avoid quasi-groups)?
- First some **notations / basics**:
 - In graph clustering we had: A graph clustering $\mathbf{C}=\{C_1, C_2, \dots, C_K\}$ is a partition of V into non-empty subsets C_k
 - Now: **clustering** $\mathcal{C} : \mathcal{X} \rightarrow \mathcal{I}$: mapping of a metric value space \mathcal{X} to a set of cluster indices \mathcal{I}
 - Clusterings can be:
 - **exclusive or non-exclusive**
 - **crisp or fuzzy**
 - **hierarchical or non-hierarchical**



Finding Clusters in Profiles

- **Exclusive** \rightarrow **non overlapping clusters**; **non-exclusive** \rightarrow overlapping clusters
- **Hierarchical** clustering \rightarrow imposes a tree structure (Dendrogram) on the C_k where an edge $C_i \rightarrow C'_j$ implies $C_i \subset C'_j$;
- **Crisp** clusterings: Conventional characteristic functions α_k for each Cluster C_k
- **Fuzzy** clusterings: fuzzy membership function α_k for each Cluster C_k

$$\alpha_k : \mathcal{X} \rightarrow \{0, 1\} \text{ with } \alpha_k(x \in \mathcal{X}) = \begin{cases} 1 & x \in C_k \\ 0 & x \notin C_k \end{cases}$$

$$\alpha_k : \mathcal{X} \rightarrow [0, 1]$$



Finding Clusters in Profiles

- **Exclusive** \rightarrow **non overlapping clusters**; **non-exclusive** \rightarrow overlapping clusters
- **Hierarchical** clustering \rightarrow imposes a tree structure (Dendrogram) on the C_k where an edge $C_i \rightarrow C'_j$ implies $C_i \subset C'_j$;
- **Crisp** clusterings: Conventional characteristic functions α_k for each Cluster C_k
- **Fuzzy** clusterings: fuzzy membership function α_k for each Cluster C_k

$$\alpha_k : \mathcal{X} \rightarrow \{0, 1\} \text{ with } \alpha_k(x \in \mathcal{X}) = \begin{cases} 1 & x \in C_k \\ 0 & x \notin C_k \end{cases}$$

$$\alpha_k : \mathcal{X} \rightarrow [0, 1]$$



Metric variant of Single / Complete link clustering

- Metric variant of **Single / Complete link clustering**: Hierarchical, crisp, non-overlapping
- **Completely analogous to graph clustering case**: Start with singletons and on each level of the dendrogram merge two clusters with minimal distance (cost)
 - **Single link**:
 - **Complete link**:

$$d(C_{k_1}, C_{k_2}) = \min_{\{n_1, n_2 | x_{n_1} \in C_{k_1} \wedge x_{n_2} \in C_{k_2}\}} \|x_{n_1} - x_{n_2}\|$$

$$d(C_{k_1}, C_{k_2}) = \max_{\{n_1, n_2 | x_{n_1} \in C_{k_1} \wedge x_{n_2} \in C_{k_2}\}} \|x_{n_1} - x_{n_2}\|$$



Metric variant of Single / Complete link clustering

- Metric variant of **Single / Complete link clustering**: Hierarchical, crisp, non-overlapping
- **Completely analogous to graph clustering case**: Start with singletons and on each level of the dendrogram merge two clusters with minimal distance (cost)

- Single link:

$$d(C_{k_1}, C_{k_2}) = \min_{\{n_1, n_2 | x_{n_1} \in C_{k_1} \wedge x_{n_2} \in C_{k_2}\}} \|x_{n_1} - x_{n_2}\|$$

- Complete link:

$$d(C_{k_1}, C_{k_2}) = \max_{\{n_1, n_2 | x_{n_1} \in C_{k_1} \wedge x_{n_2} \in C_{k_2}\}} \|x_{n_1} - x_{n_2}\|$$



K-Means Clustering

- **K-Means: Optimize intra cluster coherence**:
 - **Find prototypes** by optimizing objective function modeling intra cluster coherence as mean square error

$$J_{SQE} = \sum_{k=1}^K \sum_{\{n | x_n \in C_k\}} \|x_n - \mu_k\|^2$$

$$\frac{dJ_{SQE}}{d\mu_k} \stackrel{!}{=} 0 \quad \Longrightarrow \quad \mu^k = \frac{1}{|C_k|} \sum_{\{n | x_n \in C_k\}} x_n$$

- \rightarrow cluster prototypes are barycenters („centers of gravity“) of their clusters.



K-Means Clustering

- General idea (also valid in graph clustering): **Optimize objective function** that formalizes clustering paradigm.

- K-Means: **Optimize intra cluster coherence**:

- Describe cluster C_k by **prototype** μ_k ; prototype need not be an actual pattern (If so, algorithm works with slight modifications as well)

- Determine cluster for each pattern x_n by **nearest neighbour rule**:

$$\mathcal{C}(x_n) = k_a \leftrightarrow \|x_n - \mu_{k_a}\| = \min_i \|x_n - \mu_k\|$$



K-Means Clustering

- **K-Means: Optimize intra cluster coherence**:

- **Find prototypes** by optimizing objective function modeling intra cluster coherence as mean square error

$$J_{SQE} = \sum_{k=1}^K \sum_{\{n | x_n \in C_k\}} \|x_n - \mu_k\|^2$$

$$\frac{dJ_{SQE}}{d\mu_k} \stackrel{!}{=} 0 \quad \Longrightarrow \quad \mu^k = \frac{1}{|C_k|} \sum_{\{n | x_n \in C_k\}} x_n$$

- \rightarrow cluster prototypes are barycenters („centers of gravity“) of their clusters.



K-Means Clustering

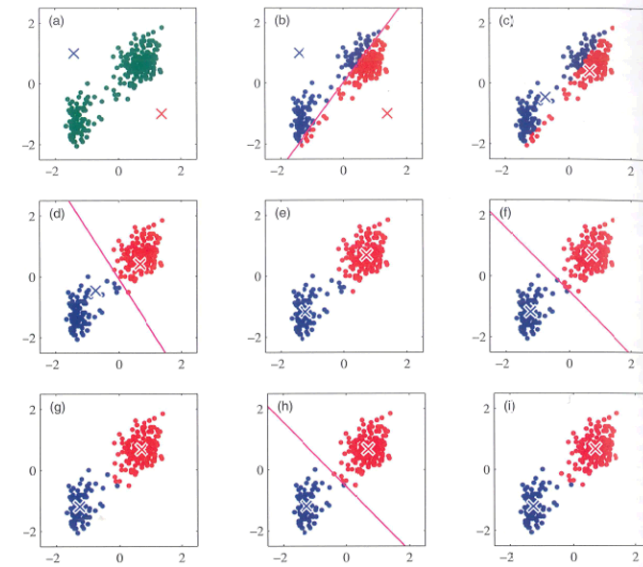
- K-Means: Optimize intra cluster coherence:
 - Find **prototypes** by optimizing objective function modeling intra cluster coherence as mean square error

$$J_{SQE} = \sum_{k=1}^K \sum_{\{n|x_n \in C_k\}} \|x_n - \mu_k\|^2$$

$$\frac{dJ_{SQE}}{d\mu_k} \stackrel{!}{=} 0 \implies \mu^k = \frac{1}{|C_k|} \sum_{\{n|x_n \in C_k\}} x_n$$

- → cluster prototypes are barycenters („centers of gravity“) of their clusters.

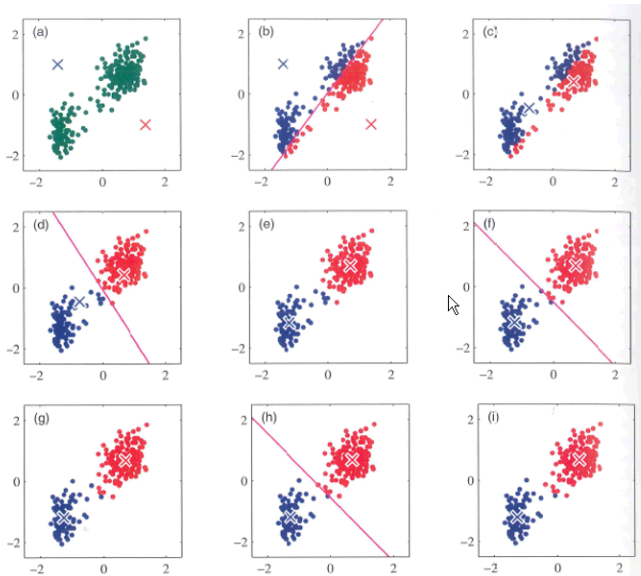
K-Means Clustering



[3]



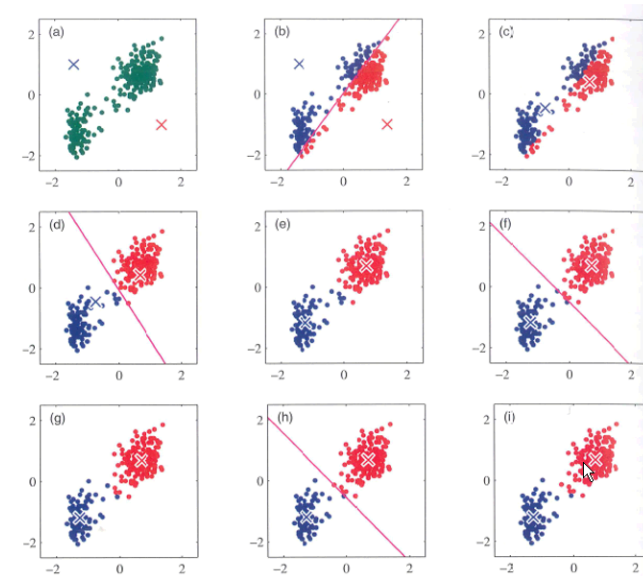
K-Means Clustering



[3]



K-Means Clustering



[3]



K-Means Clustering

- Dunn Index:

$$D = \min_{k_1 \in [1, K]} \left(\min_{k_2 \in [1, K]} \left(\frac{d_1(\mathcal{C}_{k_1}, \mathcal{C}_{k_2})}{\max_{k_3 \in [1, K]} d_2(\mathcal{C}_{k_3})} \right) \right)$$

where $d_1(\mathcal{C}_{k_1}, \mathcal{C}_{k_2})$ is the distance function between two clusters defined by

$$d_1(\mathcal{C}_{k_1}, \mathcal{C}_{k_2}) = \min_{\{(n_1, n_2) | x_{n_1} \in \mathcal{C}_{k_1} \wedge x_{n_2} \in \mathcal{C}_{k_2}\}} \|x_{n_1} - x_{n_2}\|$$

(that is the single link distance from SAHN).

The “diameter” d_2 of the clusters is defined by

$$d_2(\mathcal{C}_i) = \max_{\{(n_1, n_2) | x_{n_1} \in \mathcal{C}_i \wedge x_{n_2} \in \mathcal{C}_i\}} \|x_{n_1} - x_{n_2}\|$$



[7]

K-Means Clustering

- Dunn Index:

$$D = \min_{k_1 \in [1, K]} \left(\min_{k_2 \in [1, K]} \left(\frac{d_1(\mathcal{C}_{k_1}, \mathcal{C}_{k_2})}{\max_{k_3 \in [1, K]} d_2(\mathcal{C}_{k_3})} \right) \right)$$

where $d_1(\mathcal{C}_{k_1}, \mathcal{C}_{k_2})$ is the distance function between two clusters defined by

$$d_1(\mathcal{C}_{k_1}, \mathcal{C}_{k_2}) = \min_{\{(n_1, n_2) | x_{n_1} \in \mathcal{C}_{k_1} \wedge x_{n_2} \in \mathcal{C}_{k_2}\}} \|x_{n_1} - x_{n_2}\|$$

(that is the single link distance from SAHN).

The “diameter” d_2 of the clusters is defined by

$$d_2(\mathcal{C}_i) = \max_{\{(n_1, n_2) | x_{n_1} \in \mathcal{C}_i \wedge x_{n_2} \in \mathcal{C}_i\}} \|x_{n_1} - x_{n_2}\|$$



[7]

K-Means Clustering

- Dunn Index:

$$D = \min_{k_1 \in [1, K]} \left(\min_{k_2 \in [1, K]} \left(\frac{d_1(\mathcal{C}_{k_1}, \mathcal{C}_{k_2})}{\max_{k_3 \in [1, K]} d_2(\mathcal{C}_{k_3})} \right) \right)$$

where $d_1(\mathcal{C}_{k_1}, \mathcal{C}_{k_2})$ is the distance function between two clusters defined by

$$d_1(\mathcal{C}_{k_1}, \mathcal{C}_{k_2}) = \min_{\{(n_1, n_2) | x_{n_1} \in \mathcal{C}_{k_1} \wedge x_{n_2} \in \mathcal{C}_{k_2}\}} \|x_{n_1} - x_{n_2}\|$$

(that is the single link distance from SAHN).

The “diameter” d_2 of the clusters is defined by

$$d_2(\mathcal{C}_i) = \max_{\{(n_1, n_2) | x_{n_1} \in \mathcal{C}_i \wedge x_{n_2} \in \mathcal{C}_i\}} \|x_{n_1} - x_{n_2}\|$$



[7]

DBSCAN

- K-Means is „OK“ as cluster algorithm, but has certain **disadvantages**:
 - favors **spherical clusters**
 - **need to know K**
 - **no notion of noise**

- **Alternative → DBSCAN [4]**
(used frequently in practice):

- Idea: Two parameters: **minPt**, ϵ

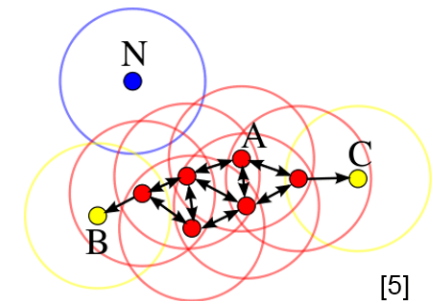
- Rough **idea**: **iterate**:

visit previously unseen pattern x:

if in ϵ -neighborhood $\{x\}$ of x: $|\{x'\}| \geq \text{minPt}$ **then**

start new cluster: include x and $\{x'\}$ and those of their ϵ -neighborhoods $\{x''\}$ that are dense enough ($|\{x''\}| \geq \text{minPt}$), etc.

else: x is noise



[5]



[7]

DBSCAN

- K-Means is „OK“ as cluster algorithm, but has certain **disadvantages**:
 - favors **spherical clusters**
 - **need to know K**
 - **no notion of noise**

- **Alternative → DBSCAN [4]**
(used frequently in practice):

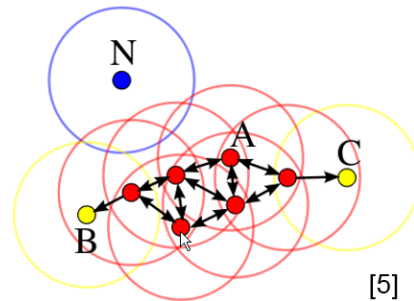
- Idea: Two parameters: **minPt, ϵ**

- Rough **idea: iterate**:

visit previously unseen pattern x:

if in ϵ -neighborhood $\{x'\}$ of x: $|\{x'\}| \geq \text{minPt}$ **then**
start new cluster: include x and $\{x'\}$ and those of their
 ϵ -neighborhoods $\{x''\}$ that are dense enough ($|\{x''\}| \geq$
minPt), etc.

else: x is noise



[5]



DBSCAN

- **Advantages** of DBSCAN:

- We do not need to know K in advance
- arbitrarily shaped clusters
- notion of noise

- **Disadvantages**:

- instead of having to know K, we need to „guess“ minPt and ϵ instead (can be a problem for high dimensional pattern spaces (→ curse of dimensionality))
- original DBSCAN has fixed (minPt, ϵ) → problems when cluster density varies



DBSCAN

- K-Means is „OK“ as cluster algorithm, but has certain **disadvantages**:
 - favors **spherical clusters**
 - **need to know K**
 - **no notion of noise**

- **Alternative → DBSCAN [4]**
(used frequently in practice):

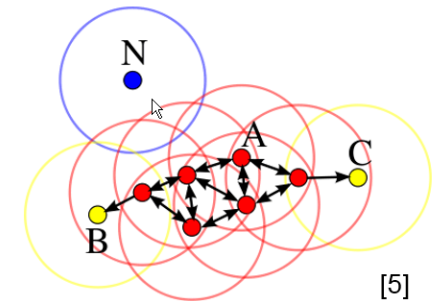
- Idea: Two parameters: **minPt, ϵ**

- Rough **idea: iterate**:

visit previously unseen pattern x:

if in ϵ -neighborhood $\{x'\}$ of x: $|\{x'\}| \geq \text{minPt}$ **then**
start new cluster: include x and $\{x'\}$ and those of their
 ϵ -neighborhoods $\{x''\}$ that are dense enough ($|\{x''\}| \geq$
minPt), etc.

else: x is noise



[5]



K-Means Clustering

- Interesting aspect: How do we **determine correct number k of clusters?**
(Same problem with graph clustering: where to cut dendrogram?)

- Answer: Compute for every k clusterings; **choose the best clustering with a cluster quality measure**

- **Cluster quality measures** for metric case: (countless variants exist in literature; for an overview: e.g. [2]) (**Objective functions** modeling clustering paradigm):

- Dunn-Index
- Entropy based indices
-



Fuzzy C-Means Clustering

- K-Means was a crisp algorithm. Now: **fuzzy variant**
- Reformulate K-Means objective function with **membership matrix** r_{nk} : Membership of pattern x_n in class C_k

$$J_{SQE} = \sum_{n=1}^N \sum_{k=1}^K r_{nk} \|x_n - \mu_k\|^2$$

- optimization criterion

$$dJ_{SQE}/d\mu_k = 0$$

- together with non-overlapping constraint

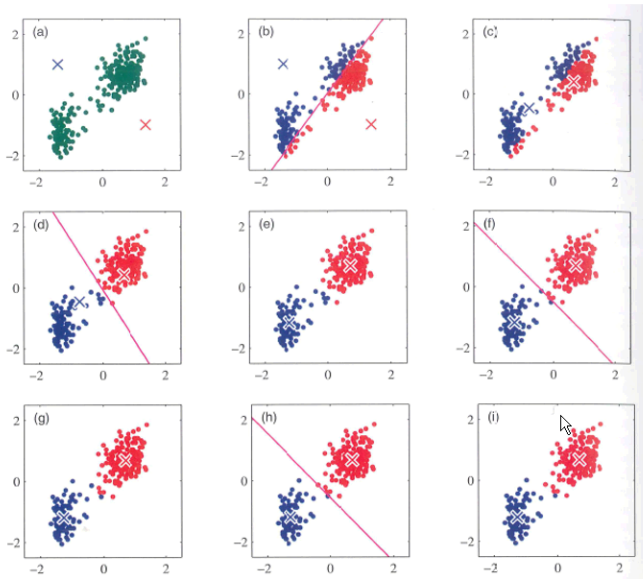
$$\forall n (\exists k (r_{nk} = 1) \wedge ((k' \neq k) \rightarrow (r_{nk'} = 0)))$$

leads to well known K-Means

$$\mu_k = \sum_{n=1}^N r_{nk} x_n / \sum_{n=1}^N r_{nk} = (1/|C_k|) \sum_{n|x_n \in C_k} x_n$$



K-Means Clustering



[3]



K-Means Clustering

Example Application: Clustering locations

- Problem: How do we **distinguish socially relevant clusters** (candidates for groups) from quasi groups?
 - **Compute clusterings over period of time**: Good candidates: clusters that appear over and over again, clusters that appear periodically
 - **Establish threshold for distance in clusters**: Human "social distance": A few meters (if groups are very small); few tens of meters (if groups are medium sized)
 - Include **velocities**: If divergent \rightarrow no group



Fuzzy C-Means Clustering

- Now **modify objective function** to:

$$J_{GSQE} = \sum_{n=1}^N \sum_{k=1}^K (r_{nk})^m \|x_n - \mu_k\|^2$$

- Exponent **m models degree of fuzziness**:
 - $m \rightarrow 1$: K-Means (crisp case);
 - $m \rightarrow \infty$: $r_{nk} \rightarrow 1/K$ (where K is the number of clusters)

- **Optimize** the obj. fct. under the **conditions**:

$$\forall x_n : \sum_{k=1}^K \alpha_k(x_n) = \sum_{k=1}^K r_{nk} = 1$$

$$\forall C_k : \sum_{n=1}^N \alpha_k(x_n) = \sum_{n=1}^N r_{nk} > 0$$



Fuzzy C-Means Clustering

- Now **modify objective function** to:

$$J_{GSQE} = \sum_{n=1}^N \sum_{k=1}^K (r_{nk})^m \|x_n - \mu_k\|^2$$

- Exponent **m models degree of fuzziness**:
 $m \rightarrow 1$: K-Means (crisp case);
 $m \rightarrow \infty$: $r_{nk} \rightarrow 1/K$ (where K is the number of clusters)

- Optimize** the obj. fct. under the **conditions**:

$$\forall x_n : \sum_{k=1}^K \alpha_k(x_n) = \sum_{k=1}^K r_{nk} = 1$$

$$\forall \mathcal{C}_k : \sum_{n=1}^N \alpha_k(x_n) = \sum_{n=1}^N r_{nk} > 0$$



Fuzzy C-Means Clustering

- Result:**

$$r_{nk} = \left(\sum_{k'=1}^K \left(\frac{\|x_n - \mu_k\|}{\|x_n - \mu_{k'}\|} \right)^{\frac{2}{m-1}} \right)^{-1} \quad (\odot)$$

$$\mu_k = \frac{\sum_{n=1}^N r_{nk}^m x_n}{\sum_{n=1}^N r_{nk}} \quad (\odot \odot)$$

- the result assumes that no patterns and prototypes coincide

$$\forall n, k : \|x_n - \mu_k\| \neq 0$$

if they do coincide, set $r_{nk} = 1$ for $x_n = \mu_k$ and $r_{nk} = 0$ for $x_n \neq \mu_k$



Fuzzy C-Means Clustering

- Now **modify objective function** to:

$$J_{GSQE} = \sum_{n=1}^N \sum_{k=1}^K (r_{nk})^m \|x_n - \mu_k\|^2$$

- Exponent **m models degree of fuzziness**:
 $m \rightarrow 1$: K-Means (crisp case);
 $m \rightarrow \infty$: $r_{nk} \rightarrow 1/K$ (where K is the number of clusters)

- Optimize** the obj. fct. under the **conditions**:

$$\forall x_n : \sum_{k=1}^K \alpha_k(x_n) = \sum_{k=1}^K r_{nk} = 1$$

$$\forall \mathcal{C}_k : \sum_{n=1}^N \alpha_k(x_n) = \sum_{n=1}^N r_{nk} > 0$$



Fuzzy C-Means Clustering

- Limit $m \rightarrow \infty$ gives:

$$r_{nk} \xrightarrow{m \rightarrow \infty} \frac{1}{\sum_{k'=1}^K 1} = \frac{1}{K}$$

- Limit $m \rightarrow 1$ we get the nearest neighbor rule (K-Means) because:

$$r_{nk} = 1 / \left(\left(\sum_{k' \neq k} \left(\frac{\|x_n - \mu_k\|}{\|x_n - \mu_{k'}\|} \right)^{\frac{2}{m-1}} \right) + 1 \right)$$

in the limit $m \rightarrow 1$ the first sum in the denominator becomes ∞ if

$$\|x_n - \mu_k\| \neq \min_{1 \leq k' \leq K} \|x_n - \mu_{k'}\|$$

and it becomes 0 if

$$\|x_n - \mu_k\| = \min_{1 \leq k' \leq K} \|x_n - \mu_{k'}\|$$



- Result:

$$r_{nk} = \left(\sum_{k'=1}^K \left(\frac{\|x_n - \mu_k\|}{\|x_n - \mu_{k'}\|} \right)^{\frac{2}{m-1}} \right)^{-1} \quad (\odot)$$

$$\mu_k = \frac{\sum_{n=1}^N r_{nk}^m x_n}{\sum_{n=1}^N r_{nk}} \quad (\odot \odot)$$

- the result assumes that no patterns and prototypes coincide

$$\forall n, k : \|x_n - \mu_k\| \neq 0$$

if they do coincide, set $r_{nk} = 1$ for $x_n = \mu_k$ and $r_{nk} = 0$ for $x_n \neq \mu_k$



Fuzzy C-Means Clustering

- Limit $m \rightarrow \infty$ gives:

$$r_{nk} \xrightarrow{m \rightarrow \infty} \frac{1}{\sum_{k'=1}^K 1} = \frac{1}{K}$$

- Limit $m \rightarrow 1$ we get the nearest neighbor rule (K-Means) because:

$$r_{nk} = 1 / \left(\left(\sum_{k' \neq k} \left(\frac{\|x_n - \mu_k\|}{\|x_n - \mu_{k'}\|} \right)^{\frac{2}{m-1}} \right) + 1 \right)$$

in the limit $m \rightarrow 1$ the first sum in the denominator becomes ∞ if

$$\|x_n - \mu_k\| \neq \min_{1 \leq k' \leq K} \|x_n - \mu_{k'}\|$$

and it becomes 0 if

$$\|x_n - \mu_k\| = \min_{1 \leq k' \leq K} \|x_n - \mu_{k'}\|$$



- Limit $m \rightarrow \infty$ gives:

$$r_{nk} \xrightarrow{m \rightarrow \infty} \frac{1}{\sum_{k'=1}^K 1} = \frac{1}{K}$$

- Limit $m \rightarrow 1$ we get the nearest neighbor rule (K-Means) because:

$$r_{nk} = 1 / \left(\left(\sum_{k' \neq k} \left(\frac{\|x_n - \mu_k\|}{\|x_n - \mu_{k'}\|} \right)^{\frac{2}{m-1}} \right) + 1 \right)$$

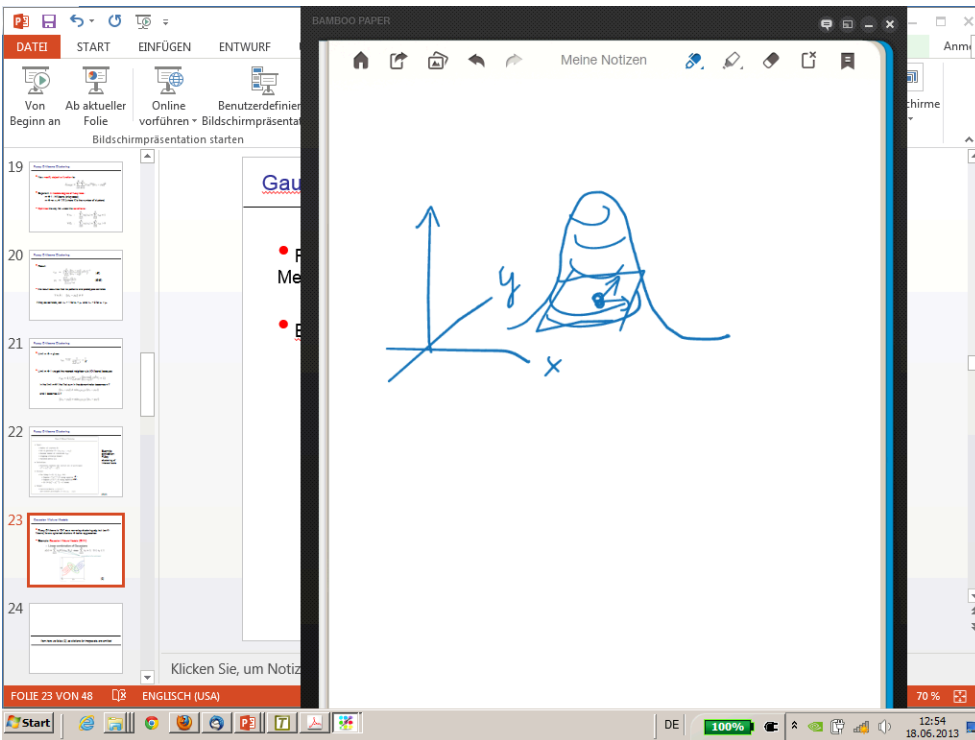
in the limit $m \rightarrow 1$ the first sum in the denominator becomes ∞ if

$$\|x_n - \mu_k\| \neq \min_{1 \leq k' \leq K} \|x_n - \mu_{k'}\|$$

and it becomes 0 if

$$\|x_n - \mu_k\| = \min_{1 \leq k' \leq K} \|x_n - \mu_{k'}\|$$





Gaussian Mixture Models

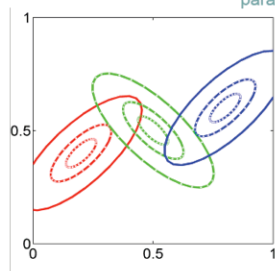
- Fuzzy C-Means is "OK" as a non-crisp clustering alg. but (as K-Means) favors spherical clusters → better approaches

- Example: **Gaussian Mixture Models (GMM)**

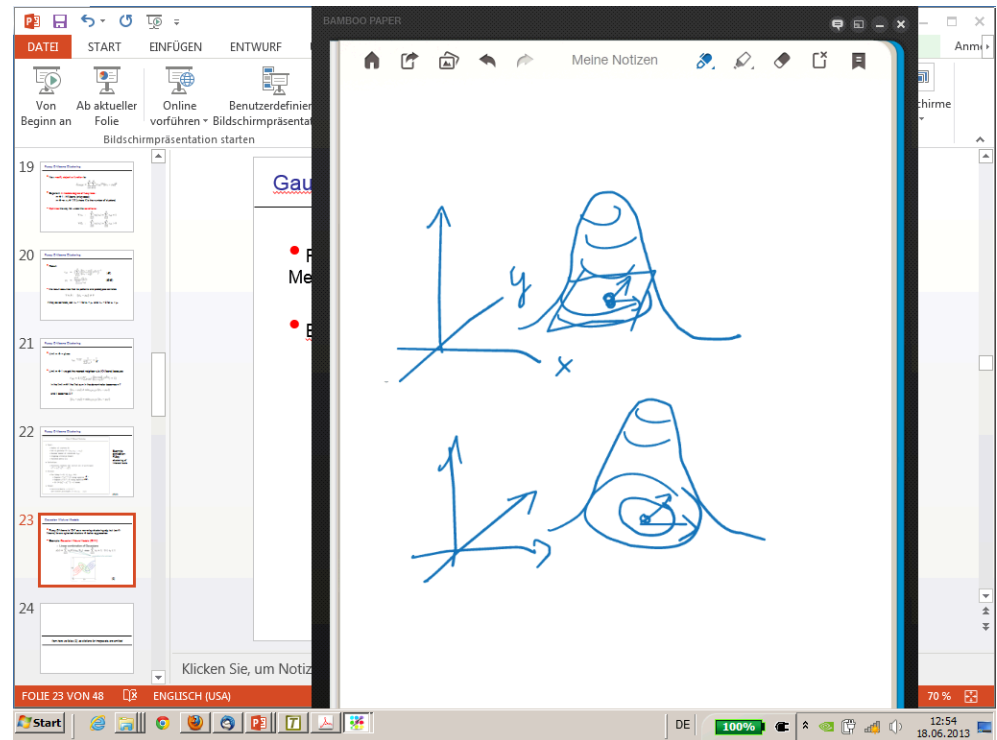
- Linear combination of Gaussians

$$p(x) = \sum_{k=1}^K \pi_k \mathcal{N}(x | \mu_k, \Sigma_k) \quad \text{where} \quad \sum_{k=1}^K \pi_k = 1, \quad 0 \leq \pi_k \leq 1$$

parameters to be estimated



[6]



Gaussian Mixture Models

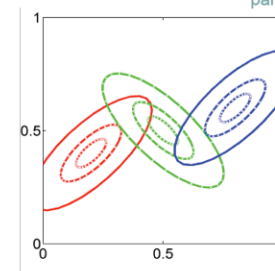
- Fuzzy C-Means is "OK" as a non-crisp clustering alg. but (as K-Means) favors spherical clusters → better approaches

- Example: **Gaussian Mixture Models (GMM)**

- Linear combination of Gaussians

$$p(x) = \sum_{k=1}^K \pi_k \mathcal{N}(x | \mu_k, \Sigma_k) \quad \text{where} \quad \sum_{k=1}^K \pi_k = 1, \quad 0 \leq \pi_k \leq 1$$

parameters to be estimated



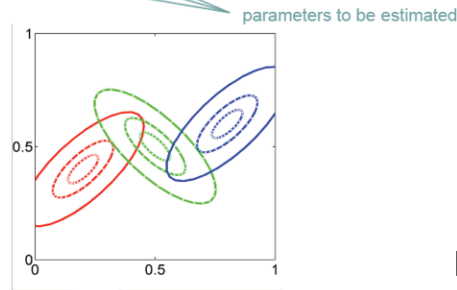
[6]

- Fuzzy C-Means is “OK” as a non-crisp clustering alg. but (as K-Means) favors spherical clusters → better approaches

- Example: **Gaussian Mixture Models (GMM)**

- Linear combination of Gaussians

$$p(x) = \sum_{k=1}^K \pi_k \mathcal{N}(x|\mu_k, \Sigma_k) \text{ where } \sum_{k=1}^K \pi_k = 1, \quad 0 \leq \pi_k \leq 1$$



[6]

Learning a **Generative Model** for data [8]:

For a distribution $p(x|\theta)$, parameterised by θ , and data $\mathcal{X} = \{x^1, \dots, x^N\}$ learning corresponds to inferring the θ that best explains the data \mathcal{X}

$$\text{Bayes theorem} \rightarrow p(\theta|\mathcal{X}) \propto p(\mathcal{X}|\theta)p(\theta)$$

- Maximum A posteriori** $\theta^{MAP} = \underset{\theta}{\operatorname{argmax}} p(\theta|\mathcal{X})$
- Maximum Likelihood** $\theta^{ML} = \underset{\theta}{\operatorname{argmax}} p(\mathcal{X}|\theta)$
 $= \underset{\theta}{\operatorname{argmax}} L(\mathcal{X}, \theta)$

Learning a **Generative Model** for data [8]:

For a distribution $p(x|\theta)$, parameterised by θ , and data $\mathcal{X} = \{x^1, \dots, x^N\}$ learning corresponds to inferring the θ that best explains the data \mathcal{X}

$$\text{Bayes theorem} \rightarrow p(\theta|\mathcal{X}) \propto p(\mathcal{X}|\theta)p(\theta)$$

- Maximum A posteriori** $\theta^{MAP} = \underset{\theta}{\operatorname{argmax}} p(\theta|\mathcal{X})$
- Maximum Likelihood** $\theta^{ML} = \underset{\theta}{\operatorname{argmax}} p(\mathcal{X}|\theta)$
 $= \underset{\theta}{\operatorname{argmax}} L(\mathcal{X}, \theta)$

Learning a **Generative Model** for data [8]:

For a distribution $p(x|\theta)$, parameterised by θ , and data $\mathcal{X} = \{x^1, \dots, x^N\}$ learning corresponds to inferring the θ that best explains the data \mathcal{X}

$$\text{Bayes theorem} \rightarrow p(\theta|\mathcal{X}) \propto p(\mathcal{X}|\theta)p(\theta)$$

- Maximum A posteriori** $\theta^{MAP} = \underset{\theta}{\operatorname{argmax}} p(\theta|\mathcal{X})$
- Maximum Likelihood** $\theta^{ML} = \underset{\theta}{\operatorname{argmax}} p(\mathcal{X}|\theta)$
 $= \underset{\theta}{\operatorname{argmax}} L(\mathcal{X}, \theta)$

Learning a **Generative Model** for data [8]:

For a distribution $p(x|\theta)$, parameterised by θ , and data $\mathcal{X} = \{x^1, \dots, x^N\}$ learning corresponds to inferring the θ that best explains the data \mathcal{X}

$$\text{Bayes theorem} \rightarrow p(\theta|\mathcal{X}) \propto p(\mathcal{X}|\theta)p(\theta)$$

- **Maximum A posteriori** $\theta^{MAP} = \underset{\theta}{\operatorname{argmax}} p(\theta|\mathcal{X})$

- **Maximum Likelihood** $\theta^{ML} = \underset{\theta}{\operatorname{argmax}} p(\mathcal{X}|\theta)$
 $= \underset{\theta}{\operatorname{argmax}} L(\mathcal{X}, \theta)$



Learning a **Generative Model** for data [8]:

For a distribution $p(x|\theta)$, parameterised by θ , and data $\mathcal{X} = \{x^1, \dots, x^N\}$ learning corresponds to inferring the θ that best explains the data \mathcal{X}

$$\text{Bayes theorem} \rightarrow p(\theta|\mathcal{X}) \propto p(\mathcal{X}|\theta)p(\theta)$$

- **Maximum A posteriori** $\theta^{MAP} = \underset{\theta}{\operatorname{argmax}} p(\theta|\mathcal{X})$

- **Maximum Likelihood** $\theta^{ML} = \underset{\theta}{\operatorname{argmax}} p(\mathcal{X}|\theta)$
 $= \underset{\theta}{\operatorname{argmax}} L(\mathcal{X}, \theta)$



Learning a **Generative Model** for data [8]:

For a distribution $p(x|\theta)$, parameterised by θ , and data $\mathcal{X} = \{x^1, \dots, x^N\}$ learning corresponds to inferring the θ that best explains the data \mathcal{X}

$$\text{Bayes theorem} \rightarrow p(\theta|\mathcal{X}) \propto p(\mathcal{X}|\theta)p(\theta)$$

- **Maximum A posteriori** $\theta^{MAP} = \underset{\theta}{\operatorname{argmax}} p(\theta|\mathcal{X})$

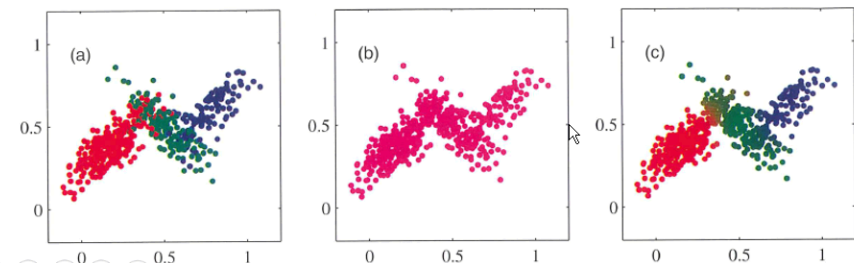
- **Maximum Likelihood** $\theta^{ML} = \underset{\theta}{\operatorname{argmax}} p(\mathcal{X}|\theta)$
 $= \underset{\theta}{\operatorname{argmax}} L(\mathcal{X}, \theta)$



- **Responsibilities**

$$\begin{aligned} \gamma(z_k) \equiv p(z_k = 1|\mathbf{x}) &= \frac{p(z_k = 1)p(\mathbf{x}|z_k = 1)}{\sum_{j=1}^K p(z_j = 1)p(\mathbf{x}|z_j = 1)} \\ &= \frac{\pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}. \end{aligned}$$

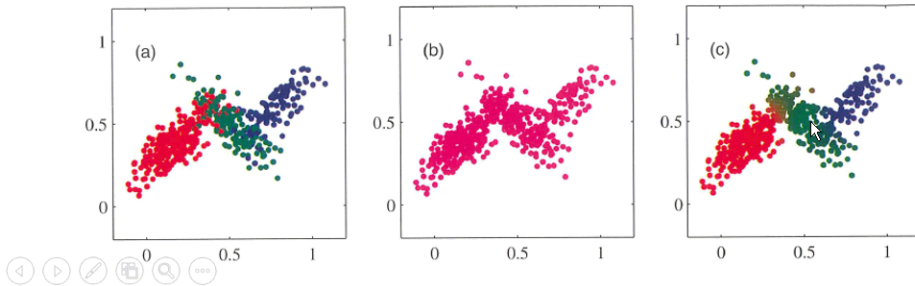
- **Example**



• Responsibilities

$$\begin{aligned} \gamma(z_k) \equiv p(z_k = 1 | \mathbf{x}) &= \frac{p(z_k = 1)p(\mathbf{x} | z_k = 1)}{\sum_{j=1}^K p(z_j = 1)p(\mathbf{x} | z_j = 1)} \\ &= \frac{\pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}. \end{aligned}$$

• Example



Maximum likelihood (one multivariate Gaussian)

$$p(\mathbf{x} | \theta) = \mathcal{N}(\mathbf{x} | \underbrace{\boldsymbol{\mu}, \boldsymbol{\Sigma}}_{\theta}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right\}$$

• Likelihood $L(\mathbf{x}, \theta) = p(\mathbf{x} | \theta)$

• Maximum likelihood $\theta_{\text{best}} = \operatorname{argmax}_{\theta} L(\mathbf{x}, \theta) = \operatorname{argmax}_{\theta} \ln L(\mathbf{x}, \theta)$

• Pattern matrix \mathbf{X} of N iid measurements (D -dim. pattern vectors \mathbf{x}),

$$\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^T$$

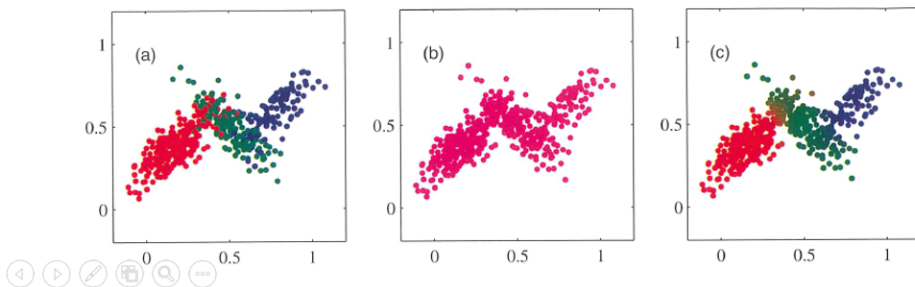
$$L(\mathbf{X}, \theta) = \prod_{i=1}^N L(\mathbf{x}_i, \theta) \quad \ln L(\mathbf{X}, \theta) = \sum_{i=1}^N \ln L(\mathbf{x}_i, \theta)$$

$$\ln L(\mathbf{X}, \theta) = \ln p(\mathbf{X} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{i=1}^N \ln N(\mathbf{x}_i | \boldsymbol{\mu}, \boldsymbol{\Sigma})$$

• Responsibilities

$$\begin{aligned} \gamma(z_k) \equiv p(z_k = 1 | \mathbf{x}) &= \frac{p(z_k = 1)p(\mathbf{x} | z_k = 1)}{\sum_{j=1}^K p(z_j = 1)p(\mathbf{x} | z_j = 1)} \\ &= \frac{\pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}. \end{aligned}$$

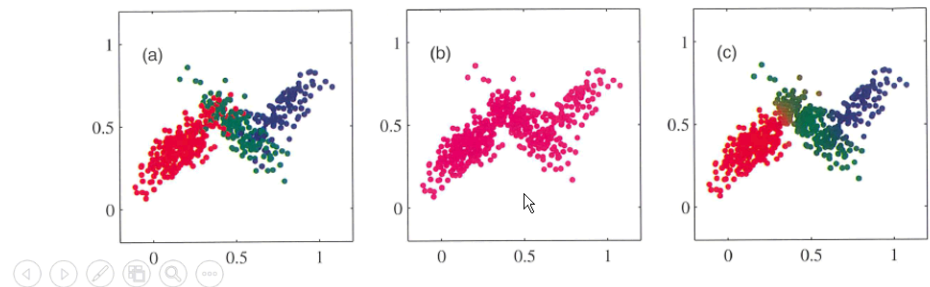
• Example



• Responsibilities

$$\begin{aligned} \gamma(z_k) \equiv p(z_k = 1 | \mathbf{x}) &= \frac{p(z_k = 1)p(\mathbf{x} | z_k = 1)}{\sum_{j=1}^K p(z_j = 1)p(\mathbf{x} | z_j = 1)} \\ &= \frac{\pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}. \end{aligned}$$

• Example



GMM-Basics

Maximum likelihood (one multivariate Gaussian)

$$p(\mathbf{x}|\theta) = \mathcal{N}(\mathbf{x}|\underbrace{\boldsymbol{\mu}, \boldsymbol{\Sigma}}_{\theta}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right\}$$

• Likelihood $L(\mathbf{x}, \theta) = p(\mathbf{x}|\theta)$

• Maximum likelihood $\theta_{\text{best}} = \operatorname{argmax}_{\theta} L(\mathbf{x}, \theta)$
 $= \operatorname{argmax}_{\theta} \ln L(\mathbf{x}, \theta)$

• Pattern matrix \mathbf{X} of N iid measurements (D -dim. pattern vectors \mathbf{x}),

$$\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^T$$

$$L(\mathbf{X}, \theta) = \prod_{i=1}^N L(\mathbf{x}_i, \theta) \quad \ln L(\mathbf{X}, \theta) = \sum_{i=1}^N \ln L(\mathbf{x}_i, \theta)$$

$$\ln L(\mathbf{X}, \theta) = \ln p(\mathbf{X} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{i=1}^N \ln N(\mathbf{x}_i | \boldsymbol{\mu}, \boldsymbol{\Sigma})$$



GMM-Basics

Maximum likelihood (one multivariate Gaussian)

$$p(\mathbf{x}|\theta) = \mathcal{N}(\mathbf{x}|\underbrace{\boldsymbol{\mu}, \boldsymbol{\Sigma}}_{\theta}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right\}$$

• Likelihood $L(\mathbf{x}, \theta) = p(\mathbf{x}|\theta)$

• Maximum likelihood $\theta_{\text{best}} = \operatorname{argmax}_{\theta} L(\mathbf{x}, \theta)$
 $= \operatorname{argmax}_{\theta} \ln L(\mathbf{x}, \theta)$

• Pattern matrix \mathbf{X} of N iid measurements (D -dim. pattern vectors \mathbf{x}),

$$\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^T$$

$$L(\mathbf{X}, \theta) = \prod_{i=1}^N L(\mathbf{x}_i, \theta) \quad \ln L(\mathbf{X}, \theta) = \sum_{i=1}^N \ln L(\mathbf{x}_i, \theta)$$

$$\ln L(\mathbf{X}, \theta) = \ln p(\mathbf{X} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{i=1}^N \ln N(\mathbf{x}_i | \boldsymbol{\mu}, \boldsymbol{\Sigma})$$



GMM-Basics

Maximum likelihood (one multivariate Gaussian)

$$p(\mathbf{x}|\theta) = \mathcal{N}(\mathbf{x}|\underbrace{\boldsymbol{\mu}, \boldsymbol{\Sigma}}_{\theta}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right\}$$

• Likelihood $L(\mathbf{x}, \theta) = p(\mathbf{x}|\theta)$

• Maximum likelihood $\theta_{\text{best}} = \operatorname{argmax}_{\theta} L(\mathbf{x}, \theta)$
 $= \operatorname{argmax}_{\theta} \ln L(\mathbf{x}, \theta)$

• Pattern matrix \mathbf{X} of N iid measurements (D -dim. pattern vectors \mathbf{x}),

$$\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^T$$

$$L(\mathbf{X}, \theta) = \prod_{i=1}^N L(\mathbf{x}_i, \theta) \quad \ln L(\mathbf{X}, \theta) = \sum_{i=1}^N \ln L(\mathbf{x}_i, \theta)$$

$$\ln L(\mathbf{X}, \theta) = \ln p(\mathbf{X} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{i=1}^N \ln N(\mathbf{x}_i | \boldsymbol{\mu}, \boldsymbol{\Sigma})$$



GMM-Basics

Maximum likelihood (one multivariate Gaussian)

$$\ln L(\mathbf{X}, \theta) = \ln L(\mathbf{X}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) =$$

$$\ln p(\mathbf{X} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = -\frac{ND}{2} \ln(2\pi) - \frac{N}{2} \ln |\boldsymbol{\Sigma}| - \frac{1}{2} \sum_{n=1}^N (\mathbf{x}_n - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x}_n - \boldsymbol{\mu})$$

$$\theta_{\text{best}} = \operatorname{argmax}_{\theta} \ln L(\mathbf{x}, \theta) \rightarrow \left. \begin{aligned} \boldsymbol{\mu}_{\text{best}} : \frac{\partial}{\partial \boldsymbol{\mu}} \ln L(\mathbf{X}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = 0 \\ \boldsymbol{\Sigma}_{\text{best}} : \frac{\partial}{\partial \boldsymbol{\Sigma}} \ln L(\mathbf{X}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = 0 \end{aligned} \right\}$$

$$\left\{ \begin{aligned} \boldsymbol{\mu}_{\text{ML}} &= \frac{1}{N} \sum_{n=1}^N \mathbf{x}_n \\ \boldsymbol{\Sigma}_{\text{ML}} &= \frac{1}{N} \sum_{n=1}^N (\mathbf{x}_n - \boldsymbol{\mu}_{\text{ML}})(\mathbf{x}_n - \boldsymbol{\mu}_{\text{ML}})^T \end{aligned} \right.$$



GMM-Basics

Maximum likelihood (one multivariate Gaussian)

$$\ln L(\mathbf{X}, \theta) = \ln L(\mathbf{X}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) =$$

$$\ln p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = -\frac{ND}{2} \ln(2\pi) - \frac{N}{2} \ln|\boldsymbol{\Sigma}| - \frac{1}{2} \sum_{n=1}^N (\mathbf{x}_n - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x}_n - \boldsymbol{\mu})$$

$$\theta_{\text{best}} = \operatorname{argmax}_{\theta} \ln L(\mathbf{x}, \theta) \rightarrow \left. \begin{aligned} \boldsymbol{\mu}_{\text{best}} &: \frac{\partial}{\partial \boldsymbol{\mu}} \ln L(\mathbf{X}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = 0 \\ \boldsymbol{\Sigma}_{\text{best}} &: \frac{\partial}{\partial \boldsymbol{\Sigma}} \ln L(\mathbf{X}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = 0 \end{aligned} \right\}$$

$$\left\{ \begin{aligned} \boldsymbol{\mu}_{\text{ML}} &= \frac{1}{N} \sum_{n=1}^N \mathbf{x}_n \\ \boldsymbol{\Sigma}_{\text{ML}} &= \frac{1}{N} \sum_{n=1}^N (\mathbf{x}_n - \boldsymbol{\mu}_{\text{ML}})(\mathbf{x}_n - \boldsymbol{\mu}_{\text{ML}})^T \end{aligned} \right.$$



GMM-Basics

$$\bullet \text{ GMM } \quad p(\mathbf{x}|\theta) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k). \quad 0 \leq \pi_k \leq 1 \quad \sum_{k=1}^K \pi_k = 1$$

1 of K representation K -dimensional binary random variable \mathbf{z}
 $z_k \in \{0, 1\}$ and $\sum_k z_k = 1$

$$p(z_k = 1) = \pi_k$$

$$p(\mathbf{z}) = \prod_{k=1}^K \pi_k^{z_k}$$

conditional probability

$$p(\mathbf{x}|z_k = 1) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \quad p(\mathbf{x}|\mathbf{z}) = \prod_{k=1}^K \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_k}$$

$$p(\mathbf{x}) = \sum_{\mathbf{z}} p(\mathbf{z}) p(\mathbf{x}|\mathbf{z}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

$p(\mathbf{x}, \mathbf{z})$



GMM-Basics

$$\bullet \text{ GMM } \quad p(\mathbf{x}|\theta) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k). \quad 0 \leq \pi_k \leq 1 \quad \sum_{k=1}^K \pi_k = 1$$

1 of K representation

K -dimensional binary random variable \mathbf{z}

$$z_k \in \{0, 1\} \text{ and } \sum_k z_k = 1$$

$$p(z_k = 1) = \pi_k$$

$$p(\mathbf{z}) = \prod_{k=1}^K \pi_k^{z_k}$$

conditional probability

$$p(\mathbf{x}|z_k = 1) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \quad p(\mathbf{x}|\mathbf{z}) = \prod_{k=1}^K \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_k}$$

$$p(\mathbf{x}) = \sum_{\mathbf{z}} p(\mathbf{z}) p(\mathbf{x}|\mathbf{z}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

$p(\mathbf{x}, \mathbf{z})$



GMM-Basics

$$\bullet \text{ GMM } \quad p(\mathbf{x}|\theta) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k). \quad 0 \leq \pi_k \leq 1 \quad \sum_{k=1}^K \pi_k = 1$$

1 of K representation

K -dimensional binary random variable \mathbf{z}

$$z_k \in \{0, 1\} \text{ and } \sum_k z_k = 1$$

$$p(z_k = 1) = \pi_k$$

$$p(\mathbf{z}) = \prod_{k=1}^K \pi_k^{z_k}$$

remark:

If we have several observations $\mathbf{x}_1, \dots, \mathbf{x}_N$, then, because we have represented the marginal distribution in the form $p(\mathbf{x}) = \sum_{\mathbf{z}} p(\mathbf{x}, \mathbf{z})$, it follows that for every observed data point \mathbf{x}_n there is a corresponding latent variable \mathbf{z}_n .

$$p(\mathbf{x}) = \sum_{\mathbf{z}} p(\mathbf{z}) p(\mathbf{x}|\mathbf{z}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

$p(\mathbf{x}, \mathbf{z})$

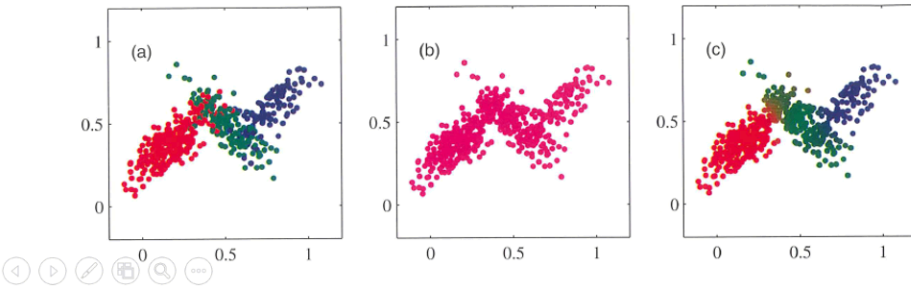


Responsibilities

$$\gamma(z_k) \equiv p(z_k = 1 | \mathbf{x}) = \frac{p(z_k = 1)p(\mathbf{x} | z_k = 1)}{\sum_{j=1}^K p(z_j = 1)p(\mathbf{x} | z_j = 1)}$$

$$= \frac{\pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

Example



Maximum likelihood (GMM)

$$\ln L(\mathbf{X}, \boldsymbol{\theta}) = \ln L(\mathbf{X}, \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \ln p(\mathbf{X} | \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

Annotations:
 - Vector of K D -dim. means $\boldsymbol{\mu}_k$
 - Vector of K $D \times D$ covariances $\boldsymbol{\Sigma}_k$

maximizing w.r.t $\boldsymbol{\pi}, \boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$ →

$$\boldsymbol{\mu}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n \quad \boldsymbol{\Sigma}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k)(\mathbf{x}_n - \boldsymbol{\mu}_k)^T$$

$$\left(N_k = \sum_{n=1}^N \gamma(z_{nk}) \right) \quad \pi_k = \frac{N_k}{N}$$

Maximum likelihood (GMM)

$$\ln L(\mathbf{X}, \boldsymbol{\theta}) = \ln L(\mathbf{X}, \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \ln p(\mathbf{X} | \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

Annotations:
 - Vector of K D -dim. means $\boldsymbol{\mu}_k$
 - Vector of K $D \times D$ covariances $\boldsymbol{\Sigma}_k$

maximizing w.r.t $\boldsymbol{\pi}, \boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$ →

$$\boldsymbol{\mu}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n \quad \boldsymbol{\Sigma}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k)(\mathbf{x}_n - \boldsymbol{\mu}_k)^T$$

$$\left(N_k = \sum_{n=1}^N \gamma(z_{nk}) \right) \quad \pi_k = \frac{N_k}{N}$$

Maximum likelihood (GMM)

$$\ln L(\mathbf{X}, \boldsymbol{\theta}) = \ln L(\mathbf{X}, \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \ln p(\mathbf{X} | \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

$$\boldsymbol{\mu}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n \quad \boldsymbol{\Sigma}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k)(\mathbf{x}_n - \boldsymbol{\mu}_k)^T$$

$$\left(N_k = \sum_{n=1}^N \gamma(z_{nk}) \right) \quad \pi_k = \frac{N_k}{N}$$

so what?! → **Problem:** Expr. depend on $\gamma(z_{nk})$ which depends on $\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}$ which depends on $\gamma(z_{nk})$ which depends on

Idea: Alternating approach (EM-algorithm):

Step t: Evaluate $\gamma(z_{nk})_{(t)}$ using $(\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})_{(t-1)}$

Evaluate $(\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})_{(t)}$ using $\gamma(z_{nk})_{(t-1)}$

Maximum likelihood (GMM)

$$\ln L(\mathbf{X}, \Theta) = \ln L(\mathbf{X}, \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \ln p(\mathbf{X} | \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

$$\boldsymbol{\mu}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n \quad \boldsymbol{\Sigma}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k)(\mathbf{x}_n - \boldsymbol{\mu}_k)^T$$

$$\left(N_k = \sum_{n=1}^N \gamma(z_{nk}) \right) \quad \pi_k = \frac{N_k}{N}$$

- so what?! → **Problem**: Expr. depend on $\gamma(z_{nk})$ which depends on $\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}$ which depends on $\gamma(z_{nk})$ which depends on

- Idea: Alternating approach (**EM-algorithm**):

Step t: Evaluate $\gamma(z_{nk})_{(t)}$ using $(\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})_{(t-1)}$

Evaluate $(\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})_{(t)}$ using $\gamma(z_{nk})_{(t-1)}$

Maximum likelihood (one multivariate Gaussian)

$$p(\mathbf{x} | \theta) = \mathcal{N}(\mathbf{x} | \underbrace{\boldsymbol{\mu}, \boldsymbol{\Sigma}}_{\theta}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right\}$$

- Likelihood** $L(\mathbf{x}, \theta) = p(\mathbf{x} | \theta)$

- Maximum likelihood** $\theta_{\text{best}} = \operatorname{argmax}_{\theta} L(\mathbf{x}, \theta)$
 $= \operatorname{argmax}_{\theta} \ln L(\mathbf{x}, \theta)$

- Pattern matrix \mathbf{X} of N iid measurements (D -dim. pattern vectors \mathbf{x}),

$$\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^T$$

$$L(\mathbf{X}, \theta) \stackrel{\text{def}}{=} \prod_{i=1}^N L(\mathbf{x}_i, \theta) \quad \ln L(\mathbf{X}, \theta) = \sum_{i=1}^N \ln L(\mathbf{x}_i, \theta)$$

$$\ln L(\mathbf{X}, \theta) = \ln p(\mathbf{X} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{i=1}^N \ln N(\mathbf{x}_i | \boldsymbol{\mu}, \boldsymbol{\Sigma})$$